

# Detection and Modulation Theory: Additional Homework

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Kay-II-3.14. Design a suitable hypothesis test for the problem

$$\left\{ \begin{array}{l} H_0 : \mathbf{x} \sim N \left( \begin{bmatrix} 0 \\ 1 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right) \\ H_1 : \mathbf{x} \sim N \left( \begin{bmatrix} 1 \\ 1 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right) \end{array} \right.$$

where  $\rho$  is given. Comment on the special case when  $\rho \rightarrow 0$ .

EL1. (After discussion by MacKay.) You are given a coin, and the task is to determine whether the coin is fair ( $H_0$ ) or not ( $H_1$ ).

a) In the first experiment you decide to flip the coin a fixed number ( $n$ ) of times and record how many heads ( $h$ ) and how many tails ( $t$ ) you got ( $n = h + t$ ). You then compute the probability  $p_e$  that *if the coin were fair, you would obtain an outcome as extreme, or more extreme, than this*. If  $p_e$  is less than a predefined, fixed “plausibility” threshold  $p_0$  (take 5% in this problem) than you determine the coin is biased ( $H_1$ ), otherwise you say it is fair ( $H_0$ ).

Suppose you decide to take  $n = 12$  (flip the coin 12 times) and you obtain the sequence hhhthhhhttht. Do you decide on  $H_0$  or  $H_1$ ?

b) In a different experiment you decide instead to flip the coin until a fixed number, say  $t$ , of tails occur. You then record how many times ( $n$ ) you had to flip the coin for this to happen. As before you then compute the

probability  $p_e$  that if the coin were fair, you would obtain an outcome that is at least as extreme as this, and decide on  $H_1$  if  $p_e$  is less than the level 5%.

Suppose you decide to take  $t = 3$  and obtained the string above, hhht hhhht hht, i.e. the 3rd tail occurred after  $n = 12$  tosses. Do you decide on  $H_0$  or  $H_1$ ?

c) Do the experiments (a) and (b) lead to the same conclusion? If not, why? Which experiment (a) or (b) do you think is most “fair”? Which one do you prefer and why?

d) Design a Bayesian test and compute the a posteriori likelihood ratio explicitly. You may make any assumptions you wish, e.g., as in the lecture.

Hints: In a) you need to consider the distribution of  $t$ , for  $n$  fixed. In b), you need to consider the distribution of  $n$ , for  $t$  fixed.

EL2. Consider the BI-AGN channel with the following signal constellation and bit-symbol mapping

$$b_1 = 0, b_2 = 0: s = -3$$

$$b_1 = 0, b_2 = 1: s = -1$$

$$b_1 = 1, b_2 = 1: s = 1$$

$$b_1 = 1, b_2 = 0: s = 3$$

(This mapping is called Gray mapping in communications.)

Suppose  $r = s + e$  where  $e \sim N(0, \sigma)$ ,  $\sigma$  known and  $b_1, b_2$  independent. Suppose  $\sigma^2 = 1$ .

a) Write up an expression for  $\text{LLR}(b_1|r)$  and  $\text{LLR}(b_2|r)$ .

b) Suppose

$$P(b_1 = 0) = P(b_1 = 1)$$

and

$$P(b_2 = 0) = P(b_2 = 1)$$

Draw  $\text{LLR}(b_1|r)$  and  $\text{LLR}(b_2|r)$  as functions of  $r$ . Carefully discuss the result.

c) Repeat the experiment for

$$P(b_1 = 0) = 1000 \cdot P(b_1 = 1)$$

and

$$P(b_2 = 0) = P(b_2 = 1)$$

How does the result for  $\text{LLR}(b_1|r)$  change? How does the result for  $\text{LLR}(b_2|r)$  change?

We assumed  $b_1, b_2$  are independent before  $r$  is observed. Are they still independent after  $r$  was observed?

Kay-I-3.9 Consider

$$r_1 = a + e_1$$

$$r_2 = a + e_2$$

and suppose  $e_1, e_2$  are jointly Gaussian with known variance  $\sigma^2$  and known correlation coefficient  $\rho$ .

Compute the CRB for the estimation of  $a$  based on  $r_1, r_2$  and sketch it as a function of  $\rho$ . Discuss the result.

EL3. Consider ML estimation (deterministic parameter vector) with “nuisance parameters”, that is, we are only interested in estimating a subset of the unknown parameters. More precisely, if the likelihood function is  $p(\mathbf{r}|\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)$  where  $\boldsymbol{\theta}_1$  and  $\boldsymbol{\theta}_2$  are unknown (vector parameters) we are only interested in estimating  $\boldsymbol{\theta}_1$ , whereas  $\boldsymbol{\theta}_2$  is an undesired nuisance parameter in which we are not particularly interested.

a) One way to deal with nuisance parameters is to form the “concentrated” ML estimate

$$\hat{\boldsymbol{\theta}}_1 = \operatorname{argmax}_{\boldsymbol{\theta}_1} \bar{p}(\mathbf{r}|\boldsymbol{\theta}_1)$$

where

$$\bar{p}(\mathbf{r}|\boldsymbol{\theta}_1) \triangleq \max_{\boldsymbol{\theta}_2} p(\mathbf{r}|\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)$$

Consider the partially linear model

$$\mathbf{r} = \mathbf{A}(\boldsymbol{\theta}_1)\boldsymbol{\theta}_2 + \mathbf{e}$$

where  $\mathbf{e} \sim N(\mathbf{0}, \mathbf{Q})$  for some positive definite  $\mathbf{Q}$ . Here  $\mathbf{A}(\boldsymbol{\theta}_1)$  indicates that  $\mathbf{A}$  is a nonlinear matrix function of  $\boldsymbol{\theta}_1$ . Derive the concentrated ML estimate of  $\boldsymbol{\theta}_1$ . You can assume that  $\mathbf{A}(\boldsymbol{\theta}_1)$  has full column rank for all  $\boldsymbol{\theta}_1$ .

b) As alternatives to “concentrating” the likelihood function (forming  $\bar{p}(\cdot)$ ), an alternative approach is to consider  $\boldsymbol{\theta}_2$  random and eliminate it by forming

$$\hat{\boldsymbol{\theta}}_1 = \operatorname{argmax}_{\boldsymbol{\theta}_1} \bar{p}(\mathbf{r}|\boldsymbol{\theta}_1)$$

where

$$\bar{p}(\mathbf{r}|\boldsymbol{\theta}_1) \triangleq E_{\boldsymbol{\theta}_2}[p(\mathbf{r}|\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)] = \int p(\mathbf{r}|\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)p(\boldsymbol{\theta}_2)d\boldsymbol{\theta}_2$$

Consider again the model  $\mathbf{r} = \mathbf{A}(\boldsymbol{\theta}_1)\boldsymbol{\theta}_2 + \mathbf{e}$  and suppose it is known a priori that  $\boldsymbol{\theta}_2 \sim N(\boldsymbol{\mu}, \boldsymbol{\Psi})$  for some positive definite  $\boldsymbol{\Psi}$ . What is the resulting estimate of  $\boldsymbol{\theta}$ ?

c) Comment on how the differences/similarities between the result in (a) and (b).

EL4. Consider the scalar Gaussian model where  $r_i$  i.i.d.

$$r_i \sim N(\mu(\theta), \sigma(\theta))$$

for  $i = 1, \dots, N$ . Derive the scalar version of Slepian-Bang's formula:

$$I(\theta) = \frac{N}{\sigma^2(\theta)} \left( \frac{d\mu(\theta)}{d\theta} \right)^2 + \frac{N}{2} \frac{1}{\sigma^4(\theta)} \left( \frac{d\sigma^2(\theta)}{d\theta} \right)^2$$

EL5. Prove the formulas for GLRT under linear Gaussian model presented in class (left column only, i.e. the case  $\boldsymbol{\theta} = \mathbf{0}$  vs  $\boldsymbol{\theta} \neq \mathbf{0}$ ).

EL6. Consider

$$H_0 : \begin{cases} x_1^i = \alpha \cos(\phi) + e_1^i \\ x_1^q = \alpha \sin(\phi) + e_1^q \\ x_2^i = \alpha \cos(\phi) + e_2^i \\ x_2^q = \alpha \sin(\phi) + e_2^q \end{cases}$$

$$H_1 : \begin{cases} x_1^i = \alpha \cos(\phi) + e_1^i \\ x_1^q = \alpha \sin(\phi) + e_1^q \\ x_2^i = -\alpha \cos(\phi) + e_2^i \\ x_2^q = -\alpha \sin(\phi) + e_2^q \end{cases}$$

You want to discriminate between  $H_0$  and  $H_1$  based on  $x_1^i, x_1^q, x_2^i, x_2^q$ . The noises  $e_x^\times$  are i.i.d.  $N(0, \sigma)$  where  $\sigma$  is known. The gain  $\alpha$  and phase  $\phi$  are unknown.

a) Explain why the problem is a simple model for demodulation of differential BPSK transmitted over a narrowband channel with amplitude  $\alpha$  and phase  $\phi$ , where  $\phi$  incorporates the phase of the previous symbol.

b) Suppose  $\phi$  is unknown, uniformly distributed over  $[-\pi, \pi]$  but  $\alpha$  is known. Construct the optimal Bayesian test.

c) Under the assumptions made in (b), compute

$$\log \left( \frac{P(H_1 | \mathbf{x}_{\times})}{P(H_0 | \mathbf{x}_{\times})} \right)$$

What assumptions do you need to compute this quantity? Why is this quantity important?

d) (optional) Comment on the case when neither  $\phi$  nor  $\alpha$  is known and suggest a decision rule.

Hint 1. In the calculations it can be helpful to work with complex numbers. However, there is no need to work with complex random numbers.

Hint 2. You may not be able evaluate all integrals in closed form. (Use Bessel functions if you want.)

EL7. (Expansion in discrete time.)

Consider a series of samples  $\{x_k = x(kT)\}_{k=-N}^N$  of a zero mean stationary continuous-time random process  $x(t)$ . Let  $\mathbf{x} = [x_{-N}, \dots, x_N]^T$ ,  $r[k] = E[x_n x_{n-k}]$  and let  $\mathbf{R} = E[\mathbf{x}\mathbf{x}^T]$  (i.e.,  $R_{ij} = r[i - j]$ ).

Let

$$Y(\omega) = \frac{1}{\sqrt{2N+1}} \sum_{k=-N}^N x_k e^{-j\omega k}$$

be the Fourier transform of  $x_k$ .

a) Compute an expression for  $E[Y(2\pi n/(2N+1))Y^*(2\pi k/(2N+1))]$ . Consider the limit when  $N \rightarrow \infty$ . Show that asymptotically  $Y(2\pi n/(2N+1))$ ,  $Y(2\pi k/(2N+1))$  are uncorrelated for  $k \neq n$  and compute their variance. Explain how their variance relates to the Fourier transform of  $r[n]$ .

For this you can assume covariance function  $r[k]$  decays exponentially fast, i.e. there is a positive constant  $\alpha$  such that  $|r[k]| = O(e^{-\alpha|k|})$

b) Argue (heuristically) that for large  $N$ , the eigenvectors to  $\mathbf{R}$  are Vandermonde vectors of the form  $[e^{-jkN2\pi/(2N+1)}, \dots, 1, \dots, e^{jkN2\pi/(2N+1)}]^T$  and the eigenvalues are given by taking uniform samples of the Fourier transform of  $r[n]$ . (A rigorous treatment can be found in R. Gray's tutorial, NOW publishers 2006.)

c) Perform an empirical study as follows. Consider a bandlimited process with covariance function  $r(t) = \sin(\omega_0 t)/(\omega_0 t)$ . Plot the eigenvalue distribution of  $\mathbf{R}$  corresponding to sampling frequency  $1/T$  for a suitably large  $N$  and some different values of  $\omega_0$ .

How many eigenvalues of  $\mathbf{R}$  are significant and how does the answer depend on  $N$ ,  $T$  and  $\omega_0$ ? How many degrees of freedom (approximately) does the sampled signal have? How do these observations relate to what we know from the sampling theorem?

EL8. Prove the formula claimed in class

$$\left(\frac{N_0}{2}\mathbf{I} + \mathbf{H}\mathbf{H}^T\right)^{-1} = \frac{2}{N_0} \left(\mathbf{I} - \mathbf{H} \left(\frac{N_0}{2}\mathbf{I} + \mathbf{H}^T\mathbf{H}\right)^{-1} \mathbf{H}^T\right) \rightarrow \frac{2}{N_0} \Pi_{\mathbf{H}}^\perp$$

(hint: use the matrix inversion lemma).