

Derivation of Cramér-Rao Bound for Unconditional Maximum Likelihood Direction of Arrival Estimation

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These notes give a detailed proof of the Cramér-Rao bound for UMA:

$$\mathbf{B}_u = \frac{\sigma}{2N} \left[\text{Re} \left\{ \mathbf{D}(\boldsymbol{\theta}_0)^H \mathbf{P}_{\mathbf{A}(\boldsymbol{\theta}_0)}^\perp \mathbf{D}(\boldsymbol{\theta}_0) \odot (\mathbf{S} \mathbf{A}(\boldsymbol{\theta}_0) \mathbf{R}(\boldsymbol{\theta}_0)^{-1} \mathbf{A}(\boldsymbol{\theta}_0) \mathbf{S}) \right\} \right]^{-1}. \quad (1)$$

In addition to that, we give show how the concentrated log-likelihood function can be used for deriving the CRB under conditional model assumption. Also, the relation

$$\frac{\mathbf{C}_{\text{cml}}}{N} \succeq \frac{\mathbf{C}_{\text{uml}}}{N} = \mathbf{B}_u$$

is proved.

I. NOTATION AND USEFUL PROPERTIES

The following notation and properties will be given without further explanation.

- $(\cdot)^H$: Hermitian (conjugate) transpose
- $(\cdot)^T$: Transpose
- $(\cdot)^*$: Conjugate
- $\text{tr}\{\cdot\}$: Trace
- $\|\cdot\|_F$: Frobenius norm
- $\mathbf{X}^\dagger = (\mathbf{X}^H \mathbf{X})^{-1} \mathbf{X}^H$: Pseudo inverse of the matrix \mathbf{X}
- $\delta(s, t) = \begin{cases} 1, & s = t \\ 0, & s \neq t \end{cases}$: The Kronecker delta function
- $\mathbf{C} = \mathbf{A} \odot \mathbf{B}$: Hadamard product, $c_{ij} = a_{ij} b_{ij}$
- $\mathbb{E}\{\cdot\}$: Expected value of a random variable.

A. Useful Trace Properties

Assume that \mathbf{B} and \mathbf{C} are matrices of appropriate dimensions, then $\text{tr}\{\mathbf{BC}\} = \text{tr}\{\mathbf{CB}\}$. Also, when \mathbf{B} and \mathbf{C} are square matrices with \mathbf{B} Hermitian we have

$$\text{tr}\{\mathbf{B}(\mathbf{C} + \mathbf{C}^H)\} = \text{tr}\{\mathbf{BC}\} + \text{tr}\{\mathbf{BC}^H\} = \text{tr}\{\mathbf{BC}\} + \text{tr}\{\mathbf{BC}\}^H = 2 \text{Re}\{\text{tr}\{\mathbf{BC}\}\}. \quad (2)$$

B. A Matrix Function Derivative

Let $\mathbf{X}(\alpha)$ be a matrix function of α . Then

$$\frac{d}{d\alpha} \log \det(\mathbf{X}(\alpha)) = \text{tr}\left\{\mathbf{X}(\alpha)^{-1} \frac{d}{d\alpha} \mathbf{X}(\alpha)\right\} \quad (3)$$

C. Matrix Inversion Lemma

Let \mathbf{A} , \mathbf{B} , \mathbf{C} , and \mathbf{D} be matrices of size $m \times m$, $m \times n$, $n \times n$, and $n \times m$, respectively and with \mathbf{A} and \mathbf{C} invertible. Then

$$(\mathbf{A} + \mathbf{BCD})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1} \mathbf{B}(\mathbf{C}^{-1} + \mathbf{DA}^{-1} \mathbf{B})^{-1} \mathbf{DA}^{-1}.$$

D. Expectation of Four Gaussian Random Variables

Theorem 1. Let a , b , c , and d be jointly Gaussian distributed. Then

$$\mathbb{E}\{abcd\} = \mathbb{E}\{ab\} \mathbb{E}\{cd\} + \mathbb{E}\{ac\} \mathbb{E}\{bd\} + \mathbb{E}\{ad\} \mathbb{E}\{bc\} - 2 \mathbb{E}\{a\} \mathbb{E}\{b\} \mathbb{E}\{c\} \mathbb{E}\{d\}.$$

The theorem is proved in [2]. There is the more general scenario of matrix-valued random variables considered.

E. Projection Matrices

With some sloppy notation, we define

$$\mathbf{P} \triangleq \mathbf{A} \mathbf{A}^\dagger = \mathbf{A}(\mathbf{A}^H \mathbf{A})^{-1} \mathbf{A}^H$$

which is the orthogonal projection onto the space spanned by \mathbf{A} . Since \mathbf{A} is a function of $\boldsymbol{\theta}$, \mathbf{P} is also a function of $\boldsymbol{\theta}$. We have the following properties:

- $\mathbf{P}^\perp = \mathbf{I} - \mathbf{P}$, $\mathbf{P} \mathbf{A} = \mathbf{A}$, and $\mathbf{P}^\perp \mathbf{A} = \mathbf{0}$

- $\mathbf{P}^2 = \mathbf{P}$ and $\mathbf{P}^H = \mathbf{P}$
- Assume that \mathbf{A} is a $m \times n$ matrix with full column rank. Then \mathbf{P} have n eigenvalues that are 1 and $m - n$ that are 0. Hence, $\text{tr}\{\mathbf{P}\} = n$ and $\text{tr}\{\mathbf{P}^\perp\} = m - n$.

We define $\mathbf{A}_k \triangleq \frac{\partial}{\partial \theta_k} \mathbf{A}$ and $\mathbf{P}_k \triangleq \frac{\partial}{\partial \theta_k} \mathbf{P}$. We will use the derivatives in [3, Sec. II-C].

II. EXPLICIT FORMULA FOR \mathbf{C}_{UML}

Here, we derive the error covariance matrix of the unconditional maximum likelihood estimate $\mathbf{C}_{\text{uml}} = \sqrt{N} \mathbb{E}\{(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)^H\}$. The UML estimate of the finite-dimensional parameter vector $\boldsymbol{\theta}$ is consistent when $N \rightarrow \infty$. Then, it follows that the UML estimate of $\boldsymbol{\theta}$ asymptotically achieves the CRB, \mathbf{B}_u , see [5]. That is, we get $\mathbf{B}_u = \mathbf{C}_{\text{uml}}/N$. The derivation of \mathbf{C}_{uml} is given in [1], but here it is more detailed.

We get the UML estimate $\hat{\boldsymbol{\theta}}$ of $\boldsymbol{\theta}_0$ as the minimum of

$$V_{\text{uml}}(\boldsymbol{\theta}) = \log \det \left(\mathbf{A} \hat{\mathbf{S}} \mathbf{A}^H + \hat{\sigma} \mathbf{I} \right). \quad (4)$$

Remember that \mathbf{A} , $\hat{\mathbf{S}}$, and $\hat{\sigma}$ are function of $\boldsymbol{\theta}$. From now, we will omit the subscript of $V(\boldsymbol{\theta})$. Since $\hat{\boldsymbol{\theta}}$ minimizes (4), we have

$$V'(\hat{\boldsymbol{\theta}}) = \mathbf{0},$$

where V' is the gradient of V with respect to $\boldsymbol{\theta}$. Also, V'' us defined to be the Hessian of V . A Taylor expansion around $\boldsymbol{\theta}_0$ gives

$$\mathbf{0} = V'(\boldsymbol{\theta}_0) + V''(\boldsymbol{\theta}_0)(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) + \dots. \quad (5)$$

Since $\hat{\boldsymbol{\theta}}$ converges to $\boldsymbol{\theta}_0$ as $N \rightarrow \infty$, second- and higher order terms can be neglected. Without affecting the asymptotic properties of $\hat{\boldsymbol{\theta}}$, we can replace $V''(\boldsymbol{\theta}_0)$ with

$$V_0''(\boldsymbol{\theta}_0) = \lim_{N \rightarrow \infty} V''(\boldsymbol{\theta}_0),$$

where $V_0''(\boldsymbol{\theta})$ is what we get when $\hat{\mathbf{R}} \rightarrow \mathbf{R}$ and $\hat{\sigma} \rightarrow \sigma$. Under weak conditions, the inverse of $V_0''(\boldsymbol{\theta})$ exists. Hence, from (5) we get

$$\sqrt{N}(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) = -\sqrt{N}[V_0''(\boldsymbol{\theta}_0)]^{-1}V'(\boldsymbol{\theta}_0).$$

Then it follows that

$$\mathbf{C}_{\text{uml}} = N[V_0''(\boldsymbol{\theta}_0)]^{-1} \lim_{N \rightarrow \infty} \mathbb{E}\{V'(\boldsymbol{\theta}_0)V'(\boldsymbol{\theta}_0)^T\} [V_0''(\boldsymbol{\theta}_0)]^{-1}.$$

Now, we can determine the gradient, $V'(\boldsymbol{\theta})$, and the Hessian, $V_0''(\boldsymbol{\theta})$.

A. Calculation of Gradient

Differentiation of $V(\boldsymbol{\theta})$ with respect to θ_i gives (3)

$$\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) = \text{tr} \left\{ (\mathbf{A}\widehat{\mathbf{S}}\mathbf{A}^H + \widehat{\sigma}\mathbf{I})^{-1} \frac{\partial}{\partial \theta_i} (\mathbf{P}\widehat{\mathbf{S}}\mathbf{P}) \right\} = \text{tr} \left\{ (\mathbf{A}\widehat{\mathbf{S}}\mathbf{A}^H + \widehat{\sigma}\mathbf{I})^{-1} \frac{\partial}{\partial \theta_i} (\mathbf{P}\widehat{\mathbf{R}}\mathbf{P} + \widehat{\sigma}\mathbf{P}^\perp) \right\}. \quad (6)$$

Using [3, eq. (18)] we get

$$\frac{\partial}{\partial \theta_i} (\mathbf{P}\widehat{\mathbf{R}}\mathbf{P} + \widehat{\sigma}\mathbf{P}^\perp) = \mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i + (\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i)^H + \widehat{\sigma}_i\mathbf{P}^\perp + \widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp + (\widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp)^H \quad (7)$$

$$= \mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i + (\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i)^H + \frac{1}{2} (\widehat{\sigma}_i\mathbf{P}^\perp + (\widehat{\sigma}_i\mathbf{P}^\perp)^H) + \widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp + (\widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp)^H \quad (8)$$

where $\widehat{\sigma}_i \triangleq \frac{\partial}{\partial \theta_i} \widehat{\sigma}$ is real (will be shown later).

Hence, (6) can be written as

$$\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) = 2 \text{Re} \left\{ \text{tr} \left\{ (\mathbf{A}\widehat{\mathbf{S}}\mathbf{A}^H + \widehat{\sigma}\mathbf{I})^{-1} \left(\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i + \widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp + \frac{1}{2}\widehat{\sigma}_i\mathbf{P}^\perp \right) \right\} \right\}. \quad (9)$$

Using the matrix inversion lemma, we have

$$(\mathbf{A}\widehat{\mathbf{S}}\mathbf{A}^H + \widehat{\sigma}\mathbf{I})^{-1} = \frac{1}{\widehat{\sigma}}\mathbf{I} - \frac{1}{\widehat{\sigma}}\mathbf{A}(\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{A} + \widehat{\sigma}\mathbf{I})^{-1}\widehat{\mathbf{S}}\mathbf{A}^H$$

and we get

$$\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) = 2 \text{Re} \left\{ \text{tr} \left\{ \left[\frac{1}{\widehat{\sigma}}\mathbf{I} - \frac{1}{\widehat{\sigma}}\mathbf{A}(\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{A} + \widehat{\sigma}\mathbf{I})^{-1}\widehat{\mathbf{S}}\mathbf{A}^H \right] \left[\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i + \frac{1}{2}\widehat{\sigma}_i\mathbf{P}^\perp + \widehat{\sigma}\mathbf{P}^\perp\mathbf{P}_i^\perp \right] \right\} \right\}. \quad (10)$$

Using [3, eq. (18)], we observe that

$$\text{tr} \{ \mathbf{P}^\perp\mathbf{P}_i^\perp \} = \text{tr} \left\{ \mathbf{P}^\perp(-\mathbf{A}_i\mathbf{A}^\dagger - \mathbf{A}\mathbf{A}_i^\dagger) \right\} = -\text{tr} \{ \mathbf{A}_i\mathbf{A}^\dagger\mathbf{P}^\perp \} - \text{tr} \{ \mathbf{P}^\perp\mathbf{A}\mathbf{A}_i^\dagger \} = 0 \quad (11)$$

By using (2) we get

$$\begin{aligned}
\text{tr}\{\widehat{\sigma}_i \mathbf{P}^\perp\} &= \widehat{\sigma}_i \text{tr}\{\mathbf{P}^\perp\} = \widehat{\sigma}_i(m-d) = \text{tr}\{\mathbf{P}_i^\perp \widehat{\mathbf{R}}\} = -\text{tr}\{\mathbf{P}_i \widehat{\mathbf{R}}\} \\
&= -\text{tr}\{\widehat{\mathbf{R}}(\mathbf{P}\mathbf{P}_i + (\mathbf{P}\mathbf{P}_i)^H)\} = -2 \text{Re}\left\{\text{tr}\{\widehat{\mathbf{R}}\mathbf{P}\mathbf{P}_i\}\right\} = -2 \text{Re}\left\{\text{tr}\{\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i\}\right\} \\
&\Leftrightarrow \\
0 &= \frac{1}{2}\text{tr}\{\widehat{\sigma}_i \mathbf{P}^\perp\} + \text{Re}\left\{\text{tr}\{\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i\}\right\}. \tag{12}
\end{aligned}$$

Note that (12) confirms that $\widehat{\sigma}_i$ is real. Since $\widehat{\mathbf{S}} = \widehat{\mathbf{R}} - \sigma \mathbf{I}$ we get

$$\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{A} + \widehat{\sigma}\mathbf{I} = (\mathbf{A}^H\mathbf{A})^{-1}\mathbf{A}^H(\widehat{\mathbf{R}} - \widehat{\sigma}\mathbf{I})\mathbf{A}(\mathbf{A}^H\mathbf{A})^{-1}\mathbf{A}^H\mathbf{A} + \widehat{\sigma}\mathbf{I} = (\mathbf{A}^H\mathbf{A})^{-1}\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A} \tag{13}$$

By using (12), (13), and $\mathbf{P}^\perp\mathbf{A} = 0$, we write (10) as

$$\begin{aligned}
\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) &= -2 \text{Re}\left\{\text{tr}\left\{\frac{1}{\widehat{\sigma}}\mathbf{A}(\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{A} + \widehat{\sigma}\mathbf{I})^{-1}\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i\right\}\right\} \\
&= -\frac{2}{\widehat{\sigma}} \text{Re}\left\{\text{tr}\left\{\mathbf{A}((\mathbf{A}^H\mathbf{A})^{-1}\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1}\widehat{\mathbf{S}}\mathbf{A}^H\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}_i\right\}\right\} \tag{14}
\end{aligned}$$

Again, using [3, eq. (18)] gives

$$\mathbf{P}_i\mathbf{A} = (\mathbf{P}^\perp\mathbf{A}_i\mathbf{A}^\dagger + (\mathbf{P}^\perp\mathbf{A}_i\mathbf{A}^\dagger)^H)\mathbf{A} = \mathbf{P}^\perp\mathbf{A}_i + (\mathbf{A}_i\mathbf{A}^\dagger)^H\mathbf{P}^\perp\mathbf{A} = \mathbf{P}^\perp\mathbf{A}_i. \tag{15}$$

By inserting (15) and [4, eq. (14)] into (14) we get

$$\begin{aligned}
\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) &= -\frac{2}{\widehat{\sigma}} \text{Re}\left\{\text{tr}\left\{(\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1}\mathbf{A}^H\mathbf{A}\mathbf{A}^\dagger(\widehat{\mathbf{S}} - \widehat{\sigma}\mathbf{I})\mathbf{A}^{\dagger H}\mathbf{A}^H\mathbf{P}\widehat{\mathbf{R}}\mathbf{P}^\perp\mathbf{A}_i\right\}\right\} \\
&= -\frac{2}{\widehat{\sigma}} \text{Re}\left\{\text{tr}\left\{(\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1}\mathbf{A}^H(\widehat{\mathbf{R}} - \widehat{\sigma}\mathbf{I})\mathbf{A}\mathbf{A}^\dagger\widehat{\mathbf{R}}\mathbf{P}^\perp\mathbf{A}_i\right\}\right\} \\
&= -\frac{2}{\widehat{\sigma}} \text{Re}\left\{\text{tr}\left\{(\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1}(\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A} - \widehat{\sigma}\mathbf{A}^H\mathbf{A})\mathbf{A}^\dagger\widehat{\mathbf{R}}\mathbf{P}^\perp\mathbf{A}_i\right\}\right\} \\
&= -\frac{2}{\widehat{\sigma}} \text{Re}\text{tr}\left\{((\mathbf{A}^H\mathbf{A})^{-1} - \widehat{\sigma}(\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1})\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{P}^\perp\mathbf{A}_i\right\} \\
&= 2 \text{Re}\left\{\text{tr}\left\{\left((\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1} - \frac{1}{\widehat{\sigma}}(\mathbf{A}^H\mathbf{A})^{-1}\right)\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{P}^\perp\mathbf{A}_i\right\}\right\}. \tag{16}
\end{aligned}$$

We note that the only non-zero column of \mathbf{A}_i is \mathbf{d}_i . Define $\widehat{\boldsymbol{\rho}}_i^H$ to be the i :th row of

$$\left((\mathbf{A}^H\widehat{\mathbf{R}}\mathbf{A})^{-1} - \frac{1}{\widehat{\sigma}}(\mathbf{A}^H\mathbf{A})^{-1}\right)\mathbf{A}^H.$$

Hence, we can write (16) as

$$\frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) = 2 \operatorname{Re} \left\{ \hat{\boldsymbol{\rho}}_i^H \hat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \right\}. \quad (17)$$

When $N \rightarrow \infty$ we can replace $\hat{\sigma}$ and $\hat{\mathbf{R}}$ with σ and \mathbf{R} respectively. This follows from the following reasoning: We write $\hat{\mathbf{R}} = \mathbf{R} + \tilde{\mathbf{R}} = \mathbf{A}^H \mathbf{S} \mathbf{A} + \sigma \mathbf{I} + \tilde{\mathbf{R}}$. Then we have $\mathbf{A}^H \hat{\mathbf{R}} \mathbf{P}^\perp = \mathbf{A}^H \tilde{\mathbf{R}} \mathbf{P}^\perp \rightarrow \mathbf{0}$, $N \rightarrow \infty$. For the scalar case we have

$$\frac{1}{r + \tilde{r}} = \frac{1/r}{1 + \tilde{r}/r} = \frac{1}{r} \left(1 - \frac{\tilde{r}}{r} + \dots \right) \approx \frac{1}{r} - \frac{\tilde{r}}{r^2}.$$

The term $\frac{\tilde{r}}{r^2} \rightarrow 0$ as $\tilde{r} \rightarrow 0$. This can be generalized to the inverse of a matrix. Hence we get

$$\begin{aligned} \frac{\partial}{\partial \theta_i} V(\boldsymbol{\theta}) &= 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \left((\mathbf{A}^H \hat{\mathbf{R}} \mathbf{A})^{-1} - \frac{1}{\hat{\sigma}} (\mathbf{A}^H \mathbf{A})^{-1} \right) \mathbf{A}^H \hat{\mathbf{R}} \mathbf{P}^\perp \mathbf{A}_i \right\} \right\} \\ &= 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \left((\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} + \text{higher order terms} \right) \mathbf{A}^H \tilde{\mathbf{R}} \mathbf{P}^\perp \mathbf{A}_i \right\} \right\} \end{aligned}$$

The higher order terms depend linearly on $\tilde{\mathbf{R}}$. Therefore, these terms can be neglected since $\tilde{\mathbf{R}} \rightarrow \mathbf{0}$. Due to the projection matrix, we can replace $\tilde{\mathbf{R}}$ with $\hat{\mathbf{R}}$. Hence, we have

$$\boldsymbol{\rho}_i^H \text{ is the } i\text{:th row of } \left((\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right) \mathbf{A}^H. \quad (18)$$

B. Calculation of the Hessian $V_0''(\boldsymbol{\theta})$

First, we determine the Hessian of $V(\boldsymbol{\theta})$. Second, we let $N \rightarrow \infty$.

$$\begin{aligned} [V''(\boldsymbol{\theta})]_{ij} &= 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \frac{\partial}{\partial \theta_j} \left[(\mathbf{A}^H \hat{\mathbf{R}} \mathbf{A})^{-1} - \frac{1}{\hat{\sigma}} (\mathbf{A}^H \mathbf{A})^{-1} \right] \mathbf{A}^H \hat{\mathbf{R}} \mathbf{P}^\perp \mathbf{A}_i \right\} \right\} \\ &\quad + 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \left[(\mathbf{A}^H \hat{\mathbf{R}} \mathbf{A})^{-1} - \frac{1}{\hat{\sigma}} (\mathbf{A}^H \mathbf{A})^{-1} \right] \left[\mathbf{A}_j^H \hat{\mathbf{R}} \mathbf{P}^\perp \mathbf{A}_i + \mathbf{A}^H \hat{\mathbf{R}} \mathbf{P}_j^\perp \mathbf{A}_i + \mathbf{A}^H \hat{\mathbf{R}} \mathbf{P}^\perp \mathbf{A}_{ij} \right] \right\} \right\}. \end{aligned} \quad (19)$$

When $N \rightarrow \infty$, $\hat{\mathbf{R}} \rightarrow \mathbf{R}$. Since $\mathbf{R} = \mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{I}$, we have

$$\begin{aligned} \mathbf{A}^H \mathbf{R} \mathbf{P}^\perp &= (\mathbf{A}^H \mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{A}^H) \mathbf{P}^\perp = \mathbf{0}, \\ \mathbf{A}_j^H \mathbf{R} \mathbf{P}^\perp &= (\mathbf{A}_j^H \mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{A}_j^H) \mathbf{P}^\perp = \sigma \mathbf{A}_j^H \mathbf{P}^\perp. \end{aligned}$$

Hence, we get

$$[V_0''(\boldsymbol{\theta})]_{ij} = 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] [\sigma \mathbf{A}_j^H \mathbf{P}^\perp + \mathbf{A}^H \mathbf{R} \mathbf{P}_j^\perp] \mathbf{A}_i \right\} \right\}. \quad (20)$$

Moreover, using [3, eq. (18)] we have

$$\begin{aligned} \mathbf{A}^H \mathbf{R} \mathbf{P}_j^\perp &= -\mathbf{A}^H \mathbf{R} \mathbf{P}_j = -\mathbf{A}^H \mathbf{R} (\mathbf{P}_j \mathbf{P} + \mathbf{P} \mathbf{P}_j) \\ &= -\mathbf{A}^H \mathbf{R} (\mathbf{P} \mathbf{A}_j \mathbf{A}^\dagger + (\mathbf{P} \mathbf{A}_j \mathbf{A}^\dagger)^H) \mathbf{P}^\perp = -\mathbf{A}^H \mathbf{R} \mathbf{A} (\mathbf{A}^H \mathbf{A})^{-1} \mathbf{A}_j^H \mathbf{P}^\perp. \end{aligned} \quad (21)$$

Inserting (21) into (20) yields

$$[V_0''(\boldsymbol{\theta})]_{ij} = 2 \operatorname{Re} \left\{ \operatorname{tr} \left\{ \left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] [\sigma \mathbf{I} - \mathbf{A}^H \mathbf{R} \mathbf{A} (\mathbf{A}^H \mathbf{A})^{-1}] \mathbf{A}_j^H \mathbf{P}^\perp \mathbf{A}_i \right\} \right\}. \quad (22)$$

Note that the only non-zero element in $\mathbf{A}_j^H \mathbf{P}^\perp \mathbf{A}_i$ is the j, i :th. Therefore,

$$[V_0''(\boldsymbol{\theta})]_{ij} = 2 \operatorname{Re} \left\{ \left[\left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] [\sigma \mathbf{I} - \mathbf{A}^H \mathbf{R} \mathbf{A} (\mathbf{A}^H \mathbf{A})^{-1}] \right]_{i,j} \right\} \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{d}_i, \quad (23)$$

where

$$\begin{aligned} & \left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] [\sigma \mathbf{I} - \mathbf{A}^H \mathbf{R} \mathbf{A} (\mathbf{A}^H \mathbf{A})^{-1}] \\ &= \sigma (\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - 2 (\mathbf{A}^H \mathbf{A})^{-1} + \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} (\mathbf{A}^H \mathbf{R} \mathbf{A}) (\mathbf{A}^H \mathbf{A})^{-1} \\ &= \sigma (\mathbf{A}^H (\mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{I}) \mathbf{A})^{-1} - 2 (\mathbf{A}^H \mathbf{A})^{-1} + \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} (\mathbf{A}^H (\mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{I}) \mathbf{A}) (\mathbf{A}^H \mathbf{A})^{-1} \\ &= \sigma (\mathbf{A}^H \mathbf{A} \mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{A}^H \mathbf{A})^{-1} - 2 (\mathbf{A}^H \mathbf{A})^{-1} + \frac{1}{\sigma} (\mathbf{S} + \sigma (\mathbf{A}^H \mathbf{A})^{-1}) \\ &= \sigma (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I})^{-1} (\mathbf{A}^H \mathbf{A})^{-1} - (\mathbf{A}^H \mathbf{A})^{-1} + \frac{\mathbf{S}}{\sigma} = \{\text{Use matrix inversion lemma}\} \\ &= \sigma \left(\frac{\mathbf{I}}{\sigma} - \frac{\mathbf{S}}{\sigma} \left(\mathbf{I} + \frac{1}{\sigma} \mathbf{A}^H \mathbf{A} \mathbf{S} \right)^{-1} \frac{\mathbf{A}^H \mathbf{A}}{\sigma} \right) (\mathbf{A}^H \mathbf{A})^{-1} - (\mathbf{A}^H \mathbf{A})^{-1} + \frac{\mathbf{S}}{\sigma} \\ &= \frac{\mathbf{S}}{\sigma} \left(\mathbf{I} - \left(\mathbf{I} + \frac{1}{\sigma} \mathbf{A}^H \mathbf{A} \mathbf{S} \right)^{-1} \right) = \{\text{Use matrix inversion lemma}\} \\ &= \frac{\mathbf{S}}{\sigma} \left(\mathbf{I} + \frac{1}{\sigma} \mathbf{A}^H \mathbf{A} \mathbf{S} \right)^{-1} \frac{1}{\sigma} \mathbf{A}^H \mathbf{A} \mathbf{S} = \frac{\mathbf{S}}{\sigma} (\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} \mathbf{A}^H \mathbf{A} \mathbf{S}. \end{aligned}$$

We note that

$$\begin{aligned}
(\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} \mathbf{A}^H \mathbf{R} &= (\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} \mathbf{A}^H (\mathbf{A} \mathbf{S} \mathbf{A} + \sigma \mathbf{I}) \\
&= (\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} (\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I}) \mathbf{A}^H = \mathbf{A}^H \\
&\Leftrightarrow \\
(\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} \mathbf{A}^H &= \mathbf{A}^H \mathbf{R}^{-1}. \tag{24}
\end{aligned}$$

Now, we can write (23) as

$$[V_0''(\boldsymbol{\theta})]_{ij} = \frac{2}{\sigma} \operatorname{Re} \left\{ [\mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S}]_{i,j} \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{d}_i \right\}, \tag{25}$$

which finally gives us the Hessian

$$V_0''(\boldsymbol{\theta}) = \frac{2}{\sigma} \operatorname{Re} \left\{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot (\mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S})^T \right\} \tag{26}$$

III. CALCULATION OF ERROR COVARIANCE

Since the $V'(\boldsymbol{\theta})$ is linear in $\widehat{\mathbf{R}}$, the second order moments of $V'(\boldsymbol{\theta})$ are of the form $\mathbb{E}\{abcd\}$.

We define

$$\begin{aligned}
[\mathbf{C}]_{ij} &\triangleq \lim_{N \rightarrow \infty} \mathbb{E} \left\{ V'(\boldsymbol{\theta}_0) V'(\boldsymbol{\theta}_0)^T \right\} \\
&= 4 \lim_{N \rightarrow \infty} \mathbb{E} \left\{ \operatorname{Re} \left\{ \boldsymbol{\rho}_i^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \right\} \operatorname{Re} \left\{ \boldsymbol{\rho}_j^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_j \right\} \right\} \\
&= 2 \lim_{N \rightarrow \infty} \mathbb{E} \left\{ \operatorname{Re} \left\{ \boldsymbol{\rho}_i^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \boldsymbol{\rho}_j^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_j \right\} + \operatorname{Re} \left\{ \boldsymbol{\rho}_i^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \mathbf{d}_j^H \widehat{\mathbf{R}} \mathbf{P}^\perp \boldsymbol{\rho}_j \right\} \right\}.
\end{aligned}$$

Note that

$$\widehat{\mathbf{R}} \mathbf{P}^\perp = \frac{1}{N} \sum_{t=1}^N \mathbf{y}(t) \mathbf{y}^H(t) \mathbf{P}^\perp = \frac{1}{N} \sum_{t=1}^N \mathbf{y}(t) (\mathbf{x}(t)^H \mathbf{A}^H + \mathbf{w}^H(t)) \mathbf{P}^\perp = \frac{1}{N} \sum_{t=1}^N \mathbf{y}(t) \mathbf{w}^H(t)$$

and $\mathbb{E}\{\mathbf{y}(t) \mathbf{w}^H(t)\} = \mathbb{E}\{(\mathbf{A} \mathbf{x}(t) + \mathbf{w}(t)) \mathbf{w}^H(t)\} = \sigma \mathbf{I}$ and $\boldsymbol{\rho}_i \mathbf{P}^\perp = \mathbf{0}$. So, by using Theorem 1, (identifying $a = \boldsymbol{\rho}_i^H \mathbf{y}(t)$, $b = \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i$, $c = \boldsymbol{\rho}_j^H \mathbf{y}(s)$, and $d = \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j$) we get

$$\begin{aligned}
\mathbb{E}\left\{\boldsymbol{\rho}_i^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \boldsymbol{\rho}_j^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_j\right\} &= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \boldsymbol{\rho}_j^H \mathbf{y}(s) \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \\
&= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \left(\mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \mathbb{E}\left\{\boldsymbol{\rho}_j^H \mathbf{y}(s) \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \right. \\
&\quad + \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \boldsymbol{\rho}_j^H \mathbf{y}(s)\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \\
&\quad + \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \boldsymbol{\rho}_j^H \mathbf{y}(s)\right\} \\
&\quad \left. - 2 \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t)\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \mathbb{E}\left\{\boldsymbol{\rho}_j^H \mathbf{y}(s)\right\} \mathbb{E}\left\{\mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \right) \\
&= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \left(\sigma \boldsymbol{\rho}_i^H \mathbf{P}^\perp \mathbf{d}_i \sigma \boldsymbol{\rho}_j \mathbf{P}^\perp \mathbf{d}_j \right. \\
&\quad + \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{y}^T(s) \boldsymbol{\rho}_j^*\right\} \mathbb{E}\left\{\mathbf{d}_i^T \mathbf{P}^\perp \mathbf{w}^*(t) \mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \\
&\quad + \sigma \delta(t, s) \boldsymbol{\rho}_i^H \mathbf{P}^\perp \mathbf{d}_j \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \boldsymbol{\rho}_j^H \mathbf{y}(s)\right\} \\
&\quad \left. - 2 \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t)\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \mathbb{E}\left\{\boldsymbol{\rho}_j^H \mathbf{y}(s)\right\} \mathbb{E}\left\{\mathbf{w}^H(s) \mathbf{P}^\perp \mathbf{d}_j\right\} \right) \\
&= 0
\end{aligned}$$

and

$$\begin{aligned}
\mathbb{E}\left\{\boldsymbol{\rho}_i^H \widehat{\mathbf{R}} \mathbf{P}^\perp \mathbf{d}_i \mathbf{d}_j^H \mathbf{P}^\perp \widehat{\mathbf{R}} \boldsymbol{\rho}_j\right\} &= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s) \mathbf{y}^H(s) \boldsymbol{\rho}_j\right\} \\
&= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \left(\mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \mathbb{E}\left\{\mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s) \mathbf{y}^H(s) \boldsymbol{\rho}_j\right\} \right. \\
&\quad + \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s)\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \mathbf{y}^H(s) \boldsymbol{\rho}_j\right\} \\
&\quad + \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{y}^H(s) \boldsymbol{\rho}_j\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s)\right\} \\
&\quad \left. - 2 \mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t)\right\} \mathbb{E}\left\{\mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \mathbb{E}\left\{\mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s)\right\} \mathbb{E}\left\{\mathbf{y}^H(s) \boldsymbol{\rho}_j\right\} \right) \\
&= \frac{1}{N^2} \sum_{t=1}^N \sum_{s=1}^N \left(\mathbb{E}\left\{\boldsymbol{\rho}_i^H \mathbf{y}(t) \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{w}(s)\right\} \mathbb{E}\left\{\boldsymbol{\rho}_j^T \mathbf{y}^*(s) \mathbf{w}^H(t) \mathbf{P}^\perp \mathbf{d}_i\right\} \right. \\
&\quad \left. + \boldsymbol{\rho}_i^H \mathbf{R} \delta(t, s) \boldsymbol{\rho}_j \mathbf{d}_j^H \mathbf{P}^\perp \sigma \mathbf{I} \delta(t, s) \mathbf{P}^\perp \mathbf{d}_i \right) \\
&= \frac{\sigma}{N} \boldsymbol{\rho}_i^H \mathbf{R} \boldsymbol{\rho}_j \mathbf{d}_j^H \mathbf{P}^\perp \mathbf{d}_i
\end{aligned}$$

Note that $\rho_i^H \mathbf{R} \rho_j$ is element (i, j) of

$$\begin{aligned}
& \left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] \mathbf{A}^H \mathbf{R} \mathbf{A} \left[(\mathbf{A}^H \mathbf{R} \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] \\
&= \left[(\mathbf{A}^H (\mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{I}) \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] \mathbf{A}^H \mathbf{R} \mathbf{A} \left[(\mathbf{A}^H (\mathbf{A} \mathbf{S} \mathbf{A}^H + \sigma \mathbf{I}) \mathbf{A})^{-1} - \frac{1}{\sigma} (\mathbf{A}^H \mathbf{A})^{-1} \right] \\
&= \left[(\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I})^{-1} - \frac{\mathbf{I}}{\sigma} \right] (\mathbf{A}^H \mathbf{A})^{-1} \mathbf{A}^H (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I}) \mathbf{A} (\mathbf{A}^H \mathbf{A})^{-1} \left[(\mathbf{A}^H \mathbf{A} \mathbf{S} + \sigma \mathbf{I})^{-1} - \frac{\mathbf{I}}{\sigma} \right] \\
&= \frac{1}{\sigma^2} \mathbf{S} \mathbf{A}^H \mathbf{A} (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I})^{-1} (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I}) (\mathbf{A}^H \mathbf{A})^{-1} (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I})^{-1} \mathbf{A}^H \mathbf{A} \mathbf{S} \\
&= \frac{1}{\sigma^2} \mathbf{S} (\mathbf{S} \mathbf{A}^H \mathbf{A} + \sigma \mathbf{I})^{-1} \mathbf{A}^H \mathbf{A} \mathbf{S} = \frac{1}{\sigma^2} \mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S},
\end{aligned}$$

where the last equality follows from (24). Also, $\mathbf{d}_j^H \mathbf{P}^\perp \mathbf{d}_i$ is element (j, i) of $\mathbf{D}^H \mathbf{P}^\perp \mathbf{D}$. So we can write

$$\mathbf{C} = \frac{2}{\sigma N} \text{Re} \{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot \mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S} \}.$$

We omit the limit, since $\mathbf{C} \rightarrow \mathbf{0}$ as $N \rightarrow \infty$. Now, we get the estimation error covariance matrix as

$$\begin{aligned}
\mathbf{C}_{\text{uml}} &= N \left[\frac{2}{\sigma} \text{Re} \{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot (\mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S})^T \} \right]^{-1} \frac{2}{\sigma N} \text{Re} \{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot \mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S} \} \\
&\quad \cdot \left[\frac{2}{\sigma} \text{Re} \{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot (\mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S})^T \} \right]^{-1} \\
&= \frac{\sigma}{2} \left[\text{Re} \{ \mathbf{D}^H \mathbf{P}^\perp \mathbf{D} \odot (\mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S})^T \} \right]^{-1}.
\end{aligned}$$

Hence, we get the Cramér-Rao bound

$$\mathbf{B}_u = \frac{\mathbf{C}_{\text{uml}}}{N}.$$

IV. ASYMPTOTIC COMPARISON BETWEEN ML ESTIMATES UNDER CMA AND UMA

We show that $\mathbf{C}_{\text{cml}} \succeq \mathbf{C}_{\text{uml}}$. We have

$$\mathbf{C}_{\text{cml}} = \frac{\sigma}{2} [\text{Re} \{ \mathbf{H} \odot \mathbf{S}^T \}]^{-1} [\text{Re} \{ \mathbf{H} \odot (\mathbf{S} \mathbf{W} \mathbf{S})^T \}] [\text{Re} \{ \mathbf{H} \odot \mathbf{S}^T \}]^{-1},$$

where $\mathbf{H} \triangleq \mathbf{D}^H \mathbf{P}^\perp \mathbf{D}$ and $\mathbf{W}^{-1} = \mathbf{S} \mathbf{A}^H \mathbf{R}^{-1} \mathbf{A} \mathbf{S}$. From [6] we have

$$[\text{Re} \{ \mathbf{X} \odot \mathbf{Z} \}]^{-1} [\text{Re} \{ \mathbf{X} \odot \mathbf{Y} \}] [\text{Re} \{ \mathbf{X} \odot \mathbf{Z} \}]^{-1} \geq [\text{Re} \{ \mathbf{X} \odot \mathbf{Y} \mathbf{Z}^{-1} \mathbf{Y} \}]^{-1}, \quad (27)$$

where \mathbf{X} and \mathbf{Z} are positive definite and \mathbf{Y} is Hermitian. By using (27) and doing some identification, we can show that $\mathbf{C}_{\text{cml}} \succeq \mathbf{C}_{\text{uml}}$.

In order to prove (27) one use that [7]

$$\begin{bmatrix} \mathbf{X} & \mathbf{Y} \\ \mathbf{Y}^T & \mathbf{Z} \end{bmatrix} \succeq \mathbf{0} \Leftrightarrow \mathbf{Z} \succeq \mathbf{0}, \mathbf{X} - \mathbf{Y}\mathbf{Z}^{-1}\mathbf{Y}^T \quad (28)$$

for Hermitian \mathbf{X} and \mathbf{Z} . The left hand side of (28) is a linear matrix inequality (LMI).

REFERENCES

- [1] P. Stoica and A. Nehorai, "Performance study of conditional and unconditional direction-of-arrival estimation," *IEEE Trans. Acoust., Speech, Signal Process.*, vol. 38, no. 10, pp. 1783–1795, Oct. 1990.
- [2] P. H. Janssen and P. Stoica, "On the expectation of the product of four matrix-valued gaussian random variables," *IEEE Trans. Automat. Contr.*, vol. 33, pp. 867–870. Sept. 1988.
- [3] T. V. K. Chaitanya, "Notes for Asymptotic Analysis Examples," Mar. 2011.
- [4] S. K. Mohammed, "Conditional versus unconditional direction of arrival estimation - Key notes," Apr. 2011.
- [5] B. Ottersten and L. Ljung, "Asymptotic results for sensor array processing," in *Proc. IEEE Int. Conf. Acoust., Speech, Signal Processing*, Glasgow, Scotland, May 1989, pp. 2266–2269.
- [6] P. Stoica and A. Nehorai, "MUSIC, maximum likelihood, and Cramér-Rao bound: Further Results and Comparisons," *IEEE Trans. Acoust., Speech, Signal Process.*, vol. 38, no. 12, pp. 2140–2150, Dec. 1990.
- [7] S. Boyd, L. El Ghaoui, E. Feron, and V. Balakrishnan, *Linear Matrix Inequalities in System and Control Theory*, SIAM, 1994.