

# Notes for Asymptotic Analysis Examples

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This notes gives a quick introduction to an array processing application and a source localization application to be discussed in the lecture of asymptotic analysis of estimators.

## I. NOTATION

- $(\cdot)^*$  : Hermitian transpose
- $\text{tr}\{\mathbf{A}\}$ : trace of Matrix  $\mathbf{A}$
- $\mathcal{R}\{\mathbf{A}\}$ : range space of  $\mathbf{A}$
- $\|\mathbf{A}\|_F$ : Frobenius norm,  $\|\mathbf{A}\|_F^2 = \text{tr}\{\mathbf{A}^* \mathbf{A}\}$
- $\mathbf{A}^\dagger$ : pseudoinverse for  $\mathbf{A}$  of full rank,  $\mathbf{A}^\dagger = (\mathbf{A}^* \mathbf{A})^{-1} \mathbf{A}^*$
- $\mathbf{P}_A$ : projection matrix,  $\mathbf{P}_A = \mathbf{A} \mathbf{A}^\dagger = \mathbf{A} (\mathbf{A}^* \mathbf{A})^{-1} \mathbf{A}^*$
- $\mathbf{P}_A^\perp$ :  $\mathbf{P}_A^\perp = \mathbf{I} - \mathbf{P}_A$

## II. ARRAY PROCESSING APPLICATION

System model and the preliminary results for subspace fitting based array processing is presented in this section. The material presented below is taken mainly from [1].

- Sensor array with  $m$  elements
- $d$  emitters at different locations,  $d < m$ .
- The output of  $i$ th sensor is

$$x_i(t) = \sum_{j=1}^d a_i(\theta_j) s_j(t) + n_i(t) \quad (1)$$

where  $a_i(\theta_j)$  is a complex scalar,  $s_j(t)$  is  $j$ th emitters signal and  $n_i(t)$  is the noise at  $i$ th sensor.

- In vector form, we can write

$$\begin{aligned}\mathbf{x}(t) &= \begin{bmatrix} \mathbf{a}(\theta_1) & \cdots & \mathbf{a}(\theta_d) \end{bmatrix} \begin{bmatrix} s_1(t) \\ \vdots \\ s_d(t) \end{bmatrix} + \begin{bmatrix} n_1(t) \\ \vdots \\ n_m(t) \end{bmatrix} \\ &= \mathbf{A}(\boldsymbol{\theta}_0) \mathbf{s}(t) + \mathbf{n}(t)\end{aligned}\quad (2)$$

where  $\mathbf{a}(\theta_j) = [a_1(\theta_j), \dots, a_m(\theta_j)]^T$  is sensor response to a unit wavefront from direction  $\theta_j$ .  $\boldsymbol{\theta}_0$  is a  $d$ -dimensional parameter vector corresponding to true DOA's.

- $G = \{\mathbf{a}(\theta_j) | \theta_j \in \Theta\}$  is the array manifold. It is assumed that the parametrization of  $G$  is known.
- $G^d = \left\{ \mathbf{A}(\boldsymbol{\theta}) | \mathbf{A}(\boldsymbol{\theta}) = \begin{bmatrix} \mathbf{a}(\theta_1) & \cdots & \mathbf{a}(\theta_d) \end{bmatrix}, \theta_1 < \theta_2 < \dots < \theta_d \right\}$ .  $G^d$  is parametrized by  $\boldsymbol{\theta} = [\theta_1 \theta_2 \dots \theta_d]^T$ .
- $\mathbf{A}(\boldsymbol{\theta}) \in G^d$  has full rank ( $d$ ).
- Sampled array output can be written as

$$\mathbf{X}_N = [\mathbf{x}(1) \dots \mathbf{x}(N)] = \mathbf{A}(\boldsymbol{\theta}_0) \mathbf{S}_N + \mathbf{N}_N \quad (3)$$

where  $\mathbf{X}_N$  is  $m \times N$ ,  $\mathbf{A}(\boldsymbol{\theta}_0)$  is  $m \times d$ ,  $\mathbf{S}_N$  is  $d \times N$  and  $\mathbf{N}_N$  is  $m \times N$ .

- $\mathbb{E}[\mathbf{n}(t) \mathbf{n}^*(t)] = \sigma^2 \mathbf{I}_m$ .
- $\mathbf{S} = \mathbb{E}[\mathbf{s}(t) \mathbf{s}^*(t)] = \lim_{n \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{s}(t) \mathbf{s}^*(t)$
- $\mathbf{R} = \mathbb{E}[\mathbf{x}(t) \mathbf{x}^*(t)] = \lim_{n \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{x}(t) \mathbf{x}^*(t) = \mathbf{A}(\boldsymbol{\theta}_0) \mathbf{S} \mathbf{A}^*(\boldsymbol{\theta}_0) + \sigma^2 \mathbf{I}_m$ .
- $\mathbf{S}$  is of rank  $d' \leq d$ . Which means that the signal waveforms are linearly dependent.
- For asymptotic analysis, the array manifold vectors are assumed to be twice continuously differentiable with bounded derivatives in a neighborhood of true DOA's.
- To enable unique identification of the signal parameters, the following condition is imposed on the array response

$$\mathbf{A}(\boldsymbol{\theta}) \mathbf{T} = \mathbf{A}(\boldsymbol{\eta}) \mathbf{U} \implies \boldsymbol{\theta} = \boldsymbol{\eta}$$

For  $\mathbf{A} \in G^d$ , this condition is met if  $d < (m + d')/2$ .

- Eigendecomposition of

$$\mathbf{R} = \sum_{i=1}^m \lambda_i \mathbf{e}_i \mathbf{e}_i^* = \mathbf{E}_s \boldsymbol{\Lambda}_s \mathbf{E}_s^* + \sigma^2 \mathbf{E}_n \mathbf{E}_n^* \quad (4)$$

where  $\lambda_1 > \lambda_2 > \dots > \lambda_{d'} > \lambda_{d'+1} = \dots = \lambda_m = \sigma^2$ .  $\mathbf{E}_s$  is  $m \times d'$ ,  $\mathbf{E}_n$  is  $m \times m - d'$  and  $\boldsymbol{\Lambda}_s$  is

$d' \times d'$ .

- $\mathcal{R}(\mathbf{E}_s) \subseteq \mathcal{R}(\mathbf{A}(\boldsymbol{\theta}_0))$  with equality iff  $d' = d$ .
- The eigendecomposition of sample covariance matrix  $\hat{\mathbf{R}} = \frac{1}{N} \mathbf{X}_N \mathbf{X}_N^* = \hat{\mathbf{E}}_s \hat{\mathbf{\Lambda}}_s \hat{\mathbf{E}}_s^* + \hat{\mathbf{E}}_n \hat{\mathbf{\Lambda}}_n \hat{\mathbf{E}}_n^*$
- It is assumed that  $d$  and  $d'$  are known.

#### A. Basic Subspace fitting problem

- We can write the basic subspace fitting problem as

$$\hat{\mathbf{A}}, \hat{\mathbf{T}} = \arg \min_{\mathbf{A}, \mathbf{T}} \|\mathbf{M} - \mathbf{A}\mathbf{T}\|_F^2 \quad (5)$$

where  $\mathbf{M}$  is  $m \times q$  matrix and  $\mathbf{T}$  is a  $p \times q$  matrix. For a fixed  $\mathbf{A}$ , the minimum w.r.t  $\mathbf{T}$  is a measure of how well the range spaces of  $\mathbf{A}$  and  $\mathbf{M}$  match. The estimate of  $\boldsymbol{\theta}$  is obtained from the parameters of  $\hat{\mathbf{A}}$ .

- The problem in (5) is separable in  $\mathbf{A}$  and  $\mathbf{T}$  [2]. Using  $\hat{\mathbf{T}} = \mathbf{A}^\dagger \mathbf{M}$ , we can write

$$\begin{aligned} \hat{\mathbf{A}} &= \arg \min_{\mathbf{A}} \|\mathbf{M} - \mathbf{A}\mathbf{A}^\dagger \mathbf{M}\|_F^2 \\ &= \arg \min_{\mathbf{A}} \|\mathbf{M} - \mathbf{P}_A \mathbf{M}\|_F^2 \\ &= \arg \min_{\mathbf{A}} \|\mathbf{P}_A^\perp \mathbf{M}\|_F^2 \\ &= \arg \max_{\mathbf{A}} \|\mathbf{P}_A \mathbf{M}\|_F^2 \\ &= \arg \max_{\mathbf{A}} \text{tr} \{ \mathbf{M}^* \mathbf{P}_A \mathbf{P}_A \mathbf{M} \} \\ &\stackrel{(a)}{=} \arg \max_{\mathbf{A}} \text{tr} \{ \mathbf{P}_A \mathbf{M} \mathbf{M}^* \} \end{aligned} \quad (6)$$

where  $\mathbf{P}_A = \mathbf{A}\mathbf{A}^\dagger$  is a projection matrix that projects onto the column space of  $\mathbf{A}$ . In (a), we have used the fact that  $\mathbf{P}_A \mathbf{P}_A = \mathbf{P}_A$  and that  $\text{tr} \{ \mathbf{M}^* \mathbf{P}_A \mathbf{M} \} = \text{tr} \{ \mathbf{P}_A \mathbf{M} \mathbf{M}^* \}$ .

##### 1) Deterministic ML:

- Treat  $\{\mathbf{s}(t)\}_{t=1}^N$  as parameters.
- $p(\mathbf{X}_N | \mathbf{S}_N, \boldsymbol{\theta}) \propto \exp\left(-\frac{1}{\sigma^2} \|\mathbf{X}_N - \mathbf{A}(\boldsymbol{\theta}) \mathbf{S}_N\|_F^2\right)$  and we have

$$\max_{\mathbf{S}_N, \boldsymbol{\theta}} p(\mathbf{X}_N | \mathbf{S}_N, \boldsymbol{\theta}) = \min_{\mathbf{S}_N, \boldsymbol{\theta}} \|\mathbf{X}_N - \mathbf{A}(\boldsymbol{\theta}) \mathbf{S}_N\|_F^2$$

- With  $\mathbf{M} = \frac{1}{\sqrt{N}} \mathbf{X}_N$  and  $\mathbf{A} \in G^d$  in (5) gives deterministic ML estimate. With  $\hat{\mathbf{S}}_N = \mathbf{A}^\dagger(\boldsymbol{\theta}) \mathbf{X}_N$ , we

have equivalent problem as

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} \text{tr} \left\{ \mathbf{P}_A(\boldsymbol{\theta}) \hat{\mathbf{R}} \right\} \quad (7)$$

- (7) is similar to (6) with  $\hat{\mathbf{R}} = \mathbf{M}\mathbf{M}^*$ .

2) *MD-MUSIC*:

- When  $d' < d$ , one dimensional MUSIC algorithm is not optimal and an equivalent multi-dimensional MUSIC algorithm was proposed in [3].
- One can write the estimator using subspace fitting formulation as

$$\begin{aligned} \hat{\boldsymbol{\theta}} &= \arg \min_{\mathbf{A} \in G^{d,T}} \left\| \hat{\mathbf{E}}_s - \mathbf{A}\mathbf{T} \right\|_F^2 \\ &= \arg \max_{\mathbf{A} \in G^d} \text{tr} \left\{ \mathbf{P}_A(\boldsymbol{\theta}) \hat{\mathbf{E}}_s \hat{\mathbf{E}}_s^* \right\} \end{aligned} \quad (8)$$

- (8) is similar to (6) with  $\mathbf{M} = \hat{\mathbf{E}}_s$ .

### B. Asymptotic Results

Lemma 2 in [1] proves that the deterministic ML estimator in (7), can equivalently (in the asymptotic distribution sense) be written as

$$\hat{\boldsymbol{\theta}} = \arg \max_{\mathbf{A} \in G^d} \text{tr} \left\{ \mathbf{P}_A(\boldsymbol{\theta}) \hat{\mathbf{E}}_s (\boldsymbol{\Lambda}_s - \sigma^2 \mathbf{I}) \hat{\mathbf{E}}_s^* \right\} \quad (9)$$

Hence for both deterministic ML and MD-MUSIC cases, we can write the parameter estimate as

$$\hat{\boldsymbol{\theta}} = \arg \max_{\boldsymbol{\theta}} V(\boldsymbol{\theta}) = \arg \max_{\boldsymbol{\theta}} \text{tr} \left\{ \mathbf{P}_A(\boldsymbol{\theta}) \hat{\mathbf{E}}_s \mathbf{W} \hat{\mathbf{E}}_s^* \right\} \quad (10)$$

with  $\mathbf{W} = (\boldsymbol{\Lambda}_s - \sigma^2 \mathbf{I})$  for deterministic ML and  $\mathbf{W} = \mathbf{I}$  for MD-MUSIC. In the following discussion we consider arbitrary  $\mathbf{W}$  which is Hermitian and positive definite.

1) *Consistency*: To be discussed in the lecture.

2) *Asymptotic Distribution*: To be discussed in the lecture.

Note: The  $d'$  largest eigenvectors of  $\hat{\mathbf{R}}$  are asymptotically Gaussian distributed with means and covariances given by

$$\mathbb{E}[\hat{\mathbf{e}}_k] = \mathbf{e}_k + O\left(\frac{1}{N}\right) \quad (11)$$

$$\mathbb{E} [(\hat{\mathbf{e}}_k - \mathbb{E}[\hat{\mathbf{e}}_k]) (\hat{\mathbf{e}}_l - \mathbb{E}[\hat{\mathbf{e}}_l])^*] = \frac{\delta_{kl}}{N} \mathbf{V}^{[k]} = \frac{\delta_{kl} \lambda_k}{N} \sum_{i=1, i \neq k} \frac{\lambda_i}{(\lambda_i - \lambda_k)^2} \mathbf{e}_i \mathbf{e}_i^* + O\left(\frac{1}{N}\right) \quad (12)$$

$$\mathbb{E} [(\hat{\mathbf{e}}_k - \mathbb{E}[\hat{\mathbf{e}}_k]) (\hat{\mathbf{e}}_l - \mathbb{E}[\hat{\mathbf{e}}_l])^T] = (1 - \delta_{kl}) \frac{-\lambda_k \lambda_l}{N (\lambda_k - \lambda_l)^2} \mathbf{e}_l \mathbf{e}_k^T + O\left(\frac{1}{N}\right) \quad (13)$$

### C. Derivatives of a Projection Matrix

- $\mathbf{P}_A = \mathbf{A} \mathbf{A}^\dagger = \mathbf{A} (\mathbf{A}^* \mathbf{A})^{-1} \mathbf{A}^* = \mathbf{P}_A^*$
- $\mathbf{P}_A \mathbf{A} = \mathbf{A}$ ,  $\mathbf{P}_A \mathbf{P}_A = \mathbf{P}_A$ .
- $\mathbf{P}_A^\perp = \mathbf{I} - \mathbf{P}_A$  and  $\mathbf{P}_A^\perp \mathbf{A} = \mathbf{0}$
- In the following, we drop the subscript  $\mathbf{A}$  for  $\mathbf{P}_A$ . Note that the the derivations given below are slightly different from the ones in [1] and are taken from [2]. Starting with

$$\mathbf{P}_\eta = \frac{d\mathbf{P}}{d\theta_\eta} = \mathbf{A}_\eta \mathbf{A}^\dagger + \mathbf{A} \mathbf{A}_\eta^\dagger \quad (14)$$

and

$$\mathbf{A}_\eta = \frac{d\mathbf{A}}{d\theta_\eta} = \frac{d\mathbf{P} \mathbf{A}}{d\theta_\eta} = \mathbf{P}_\eta \mathbf{A} + \mathbf{P} \mathbf{A}_\eta \implies \mathbf{P}_\eta \mathbf{A} = \mathbf{P}^\perp \mathbf{A}_\eta \quad (15)$$

Since  $\mathbf{P} = \mathbf{A} \mathbf{A}^\dagger$ , using (15), we have

$$\mathbf{P}_\eta \mathbf{P} = \mathbf{P}_\eta \mathbf{A} \mathbf{A}^\dagger = \mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger \quad (16)$$

We also have that

$$(\mathbf{P}_\eta \mathbf{P})^* = \mathbf{P} \mathbf{P}_\eta \quad (17)$$

Now using that  $\mathbf{P} = \mathbf{P} \mathbf{P}$ , we have

$$\mathbf{P}_\eta = \frac{d\mathbf{P}}{d\theta_\eta} = \frac{d(\mathbf{P} \mathbf{P})}{d\theta_\eta} = \mathbf{P}_\eta \mathbf{P} + \mathbf{P} \mathbf{P}_\eta = \mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger + (\mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger)^* \quad (18)$$

hence from (14) and (18), we have

$$\begin{aligned} \mathbf{A}_\eta \mathbf{A}^\dagger + \mathbf{A} \mathbf{A}_\eta^\dagger &= \mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger + (\mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger)^* \\ \implies \mathbf{A} \mathbf{A}_\eta^\dagger &= \mathbf{P}^\perp \mathbf{A}_\eta \mathbf{A}^\dagger - \mathbf{A}_\eta \mathbf{A}^\dagger + \mathbf{A}^\dagger^* \mathbf{A}_\eta^* \mathbf{P}^\perp \\ &= -\mathbf{P} \mathbf{A}_\eta \mathbf{A}^\dagger + \mathbf{A}^\dagger^* \mathbf{A}_\eta^* \mathbf{P}^\perp \end{aligned}$$

$$= -\mathbf{P}\mathbf{A}_\eta\mathbf{A}^\dagger + \mathbf{A}(\mathbf{A}^*\mathbf{A})^{-1}\mathbf{A}_\eta^*\mathbf{P}^\perp \quad (19)$$

From (19), we can write (multiplying with  $\mathbf{A}^\dagger$  from the left side)

$$\begin{aligned} \mathbf{A}_\eta^\dagger &= \mathbf{A}^\dagger\mathbf{A}(\mathbf{A}^*\mathbf{A})^{-1}\mathbf{A}_\eta^*\mathbf{P}^\perp - \mathbf{A}^\dagger\mathbf{P}\mathbf{A}_\eta\mathbf{A}^\dagger \\ &= (\mathbf{A}^*\mathbf{A})^{-1}\mathbf{A}_\eta^*\mathbf{P}^\perp - \mathbf{A}^\dagger\mathbf{A}_\eta\mathbf{A}^\dagger \end{aligned} \quad (20)$$

- Note that  $\text{tr}(\mathbf{P}_\eta) = 0$ .
- From (18), we can write

$$\mathbf{P}_{\eta\xi} = \mathbf{P}_\xi^\perp\mathbf{A}_\eta\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_{\eta\xi}\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_\eta\mathbf{A}_\xi^\dagger + \left(\mathbf{P}_\xi^\perp\mathbf{A}_\eta\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_{\eta\xi}\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_\eta\mathbf{A}_\xi^\dagger\right)^* \quad (21)$$

- Noting that  $\mathbf{P}_\xi^\perp = -\mathbf{P}_\xi$  and using (18) and (20) in (21), we get

$$\begin{aligned} \mathbf{P}_{\eta\xi} &= -\mathbf{P}^\perp\mathbf{A}_\xi\mathbf{A}^\dagger\mathbf{A}_\eta\mathbf{A}^\dagger - \mathbf{A}^{\dagger*}\mathbf{A}_\xi^*\mathbf{P}^\perp\mathbf{A}_\eta\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_{\eta\xi}\mathbf{A}^\dagger + \mathbf{P}^\perp\mathbf{A}_\eta(\mathbf{A}^*\mathbf{A})^{-1}\mathbf{A}_\xi^*\mathbf{P}^\perp \\ &\quad - \mathbf{P}^\perp\mathbf{A}_\eta\mathbf{A}^\dagger\mathbf{A}_\xi\mathbf{A}^\dagger + (\dots)^* \end{aligned} \quad (22)$$

where  $(\dots)^*$  means same expression appears with Hermitian transpose.

#### D. Some tricks used in the paper

- 1)  $\text{tr}(\mathbf{ABC}) = \text{tr}(\mathbf{CAB}) = \text{tr}(\mathbf{BCA})$
- 2) Since  $\mathcal{R}(\mathbf{E}_s) \subseteq \mathcal{R}(\mathbf{A})$ ,  $\mathbf{P}_\mathbf{A}^\perp\mathbf{E}_s = \mathbf{0} \implies \mathbf{E}_s^*\mathbf{P}_\mathbf{A}^\perp = \mathbf{0}$
- 3) If  $\mathbf{B}$  is Hermitian, and  $\mathbf{A}$  has the same dimensions as  $\mathbf{B}$ , we have

$$\text{tr}\{(\mathbf{A} + \mathbf{A}^*)\mathbf{B}\} = \text{tr}\{\mathbf{AB}\} + \text{tr}\{\mathbf{A}^*\mathbf{B}\} = 2 \text{Re}[\text{tr}\{\mathbf{BA}\}] \quad (23)$$

- 4) Let  $a_k = a_{kr} + ja_{ki}$  denote a complex number with  $a_{kr}$  and  $a_{ki}$  being the real and imaginary parts of  $a_k$ . Suppose  $\{a_k\}_{k=1}^d$  and  $\{b_k\}_{k=1}^d$  are sets of complex numbers, then we have

$$2 \text{Re} \left\{ \sum_k a_k \right\} \cdot 2 \text{Re} \left\{ \sum_k b_k \right\} = 4 \left\{ \sum_k \sum_l a_{kr} b_{lr} \right\} \quad (24)$$

We can write (24) equivalently as

$$2 \text{Re} \left\{ \sum_k a_k \right\} \cdot 2 \text{Re} \left\{ \sum_k b_k \right\} = 2 \text{Re} \left\{ \sum_k \sum_l (a_k b_l + a_k^* b_l) \right\} = 4 \left\{ \sum_k \sum_l a_{kr} b_{lr} \right\} \quad (25)$$

5) When  $\mathbf{S} = \mathbb{E}[\mathbf{s}(t)\mathbf{s}^*(t)] = \lim_{n \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{s}(t)\mathbf{s}^*(t)$  is full rank ( $d$ ). Then  $\mathbf{E}_s^\dagger = \mathbf{E}_s^*$  and  $\mathbf{P}_A = \mathbf{E}_s \mathbf{E}_s^*$ .

Proof: Note that when  $\mathbf{S}$  is full rank,  $\mathcal{R}(\mathbf{A}) = \mathcal{R}(\mathbf{E}_s)$ . Where  $\mathbf{E}_s$  and  $\mathbf{A}$  are  $m \times d$  matrices. This implies that there exists a invertible  $d \times d$  matrix  $\mathbf{Z}$  such that

$$\mathbf{A} = \mathbf{E}_s \mathbf{Z} \quad (26)$$

Noting that

$$\mathbf{R} = \mathbb{E}[\mathbf{x}(t)\mathbf{x}^*(t)] = \lim_{n \rightarrow \infty} \frac{1}{N} \sum_{t=1}^N \mathbf{x}(t)\mathbf{x}^*(t) = \mathbf{A} \mathbf{S} \mathbf{A}^* + \sigma^2 \mathbf{I}_m \quad (27)$$

has eigendecomposition as

$$\mathbf{R} = \begin{bmatrix} \mathbf{E}_s & \mathbf{E}_n \end{bmatrix} \begin{bmatrix} \Lambda_s & \\ & \sigma^2 \mathbf{I} \end{bmatrix} \begin{bmatrix} \mathbf{E}_s^* \\ \mathbf{E}_n^* \end{bmatrix} = \mathbf{E}_s \Lambda_s \mathbf{E}_s^* + \sigma^2 \mathbf{E}_n \mathbf{E}_n^* \quad (28)$$

with

$$\mathbf{E}_s \mathbf{E}_s^* + \mathbf{E}_n \mathbf{E}_n^* = \mathbf{I}_m \quad (29)$$

and  $\mathbf{E}_s^* \mathbf{E}_s = \mathbf{I}_d$  and  $\mathbf{E}_n^* \mathbf{E}_n = \mathbf{I}_{m-d}$ . From which we have  $\mathbf{E}_s^\dagger = (\mathbf{E}_s^* \mathbf{E}_s)^{-1} \mathbf{E}_s^* = \mathbf{E}_s^*$ .

Now  $\mathbf{P}_A = \mathbf{A}(\mathbf{A}^* \mathbf{A})^{-1} \mathbf{A}^* = \mathbf{E}_s \mathbf{Z} (\mathbf{Z}^* \mathbf{E}_s^* \mathbf{E}_s \mathbf{Z})^{-1} \mathbf{Z}^* \mathbf{E}_s^* = \mathbf{E}_s \mathbf{E}_s^*$ .

### III. SOURCE LOCALIZATION APPLICATION

The system model for the source localization problem from [4] is presented below.

- A source  $\mathcal{S}$  is located at  $(x_0, y_0) \in \mathbb{R}^2$ .
- A set  $\mathcal{A} \triangleq \{\mathcal{A}_1, \dots, \mathcal{A}_M\}$  of  $M$  anchors are used for locating the source. Anchor  $\mathcal{A}_m$  is located at  $(x_m, y_m) \in \mathbb{R}^2$ .
- The true distance between  $\mathcal{S}$  and  $\mathcal{A}_m$  is denoted as

$$d_m = \sqrt{(x_0 - x_m)^2 + (y_0 - y_m)^2} \quad (30)$$

- $N$  independent measurements  $r_{m,n}$  of each distance  $d_m$  are available

$$r_{m,n} = d_m + e_{m,n} = \sqrt{(x_0 - x_m)^2 + (y_0 - y_m)^2} + e_{m,n}, \quad n = 1, 2, \dots, N \quad (31)$$

where  $e_{m,n}$  are measurement errors.

- We consider two source localization methods
  - Least squares (LS)
  - Squared-range least squares (SR-LS)

- LS:

$$(\hat{x}_{LS}, \hat{y}_{LS}) = \arg \min_{x,y} \sum_{n=1}^N \sum_{m=1}^N \left( r_{m,n} - \sqrt{(x - x_m)^2 + (y - y_m)^2} \right)^2 \quad (32)$$

- We can show that the problem in (32) can be equivalently be written as

$$f_{LS}(x, y) \triangleq \sum_{m=1}^M \left( r_m - \sqrt{(x - x_m)^2 + (y - y_m)^2} \right)^2 \quad (33)$$

where  $r_m \triangleq \frac{1}{N} \sum_{n=1}^N r_{m,n} = \sqrt{(x_0 - x_m)^2 + (y_0 - y_m)^2} + e_m$  and  $e_m = \frac{1}{N} \sum_{n=1}^N e_{m,n}$  is the averaged noise.

- LS is equivalent to ML if  $e_{m,n}$  are i.i.d zero mean Gaussian distributed.
- SR-LS:

$$f_{SR-LS}(x, y) \triangleq \sum_{m=1}^M \left( r_m^2 - ((x - x_m)^2 + (y - y_m)^2) \right)^2 \quad (34)$$

- For  $M = 2$ , LS and SR-LS are equivalent in all cases of practical interest.
- For  $M > 2$ , LS and SR-LS are not equivalent.
- Assuming that  $e_{m,n}$  are independent with zero mean,  $\mathbb{E}[e_{m,n}^2] = \sigma^2$ ,  $\mathbb{E}[e_{m,n}^3] = 0$ , and  $\mathbb{E}[e_{m,n}^4] = \alpha\sigma^4$  for some constant  $\alpha$ . For Gaussian case  $\alpha = 3$ .
- This results in  $\mathbb{E}(e_m) = 0$ ,  $\mathbb{E}(e_m^2) = \frac{\sigma^2}{N}$  and  $\mathbb{E}(e_m^4) = \frac{\sigma^4}{N^2} \left( 3 + \frac{\alpha-3}{N} \right)$ .
- For convenience, without loss of generality, we assume that true location of source  $\mathcal{S}$  is  $(x_0, y_0) = (0, 0)$ .
- Some notations

$$X_a(\mathcal{A}) \triangleq \sum_{m=1}^M \frac{x_m^2}{x_m^2 + y_m^2}, Y_a(\mathcal{A}) \triangleq \sum_{m=1}^M \frac{y_m^2}{x_m^2 + y_m^2}, Z_a(\mathcal{A}) \triangleq \sum_{m=1}^M \frac{x_m y_m}{x_m^2 + y_m^2}$$

$$X_b(\mathcal{A}) \triangleq \sum_{m=1}^M x_m^2, Y_b(\mathcal{A}) \triangleq \sum_{m=1}^M y_m^2, Z_b(\mathcal{A}) \triangleq \sum_{m=1}^M x_m y_m$$



$$X_c(\mathcal{A}) \triangleq \sum_{m=1}^M x_m^2 (x_m^2 + y_m^2), Y_c(\mathcal{A}) \triangleq \sum_{m=1}^M y_m^2 (x_m^2 + y_m^2), Z_c(\mathcal{A}) \triangleq \sum_{m=1}^M x_m y_m (x_m^2 + y_m^2)$$

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